# Exact Average Run Length Evaluation for an ARMAX(p,q,r) Process Running on a Modified EWMA Control Chart 

Korakoch Silpakob, Yupaporn Areepong, Saowanit Sukparungsee and Rapin Sunthornwat


#### Abstract

In this study, we apply the Fredholm-type integral equation method to derive the explicit formulas of the average run length (ARL) for an autoregressive moving average process with explanatory variables (ARMAX(p,q,r)) with exponential white noise running on a modified exponentially weighted moving average (EWMA) control chart. As a performance measure, we compared the computational times of calculating the ARL based on explicit formulas and the classical numerical integral equation (NIE) method. We found that although the ARLs using both methods were very close with an absolute percentage difference of less than $\mathbf{0 . 0 0 0 0 1 \%}$, their calculational times were less than 0.01 and 10 seconds, respectively. Furthermore, the comparison of the performances of the ARL methods for ARMAX( $\mathbf{p , q , r )}$ processes with exponential white noise by practical application for time series data comprising exchange rates and the price of energy running on modified and standard EWMA and cumulative sum (CUSUM) control charts using the relative mean index (RMI) criteria. The results show that the explicit formulas method for the ARL of the process on the modified EWMA control chart is more powerful than the CUSUM and standard EWMA control charts.


Index Terms- Autoregressive process, moving average process, explanatory variable, explicit formulas

## I. Introduction

STATISTICAL process control (SPC) is a powerful set of tools that are used to inspect, control, and improve the quality of products or services that plays an essential role in business and manufacturing sectors. Control charts are one of the key tools in SPC widely used in various fields, such

[^0]as health [1], medicine [2], and finance [3]. Shewhart [4] presented the first control chart that is still widely used for monitoring and detecting large shifts in a process mean. Later, the cumulative sum (CUSUM) control chart [5] and the standard exponentially weighted moving average (EWMA) control chart [6] were found to be more suitable for detecting small shifts in a process mean. Moreover, Khan et al. [7] modified the EWMA control chart by adding an extra constant ( $k$ ) in the last term of the modified EWMA statistic, which was further modified by [8]. The authors compared its performance with the originally modified and standard EWMA control charts in terms of the average run length (ARL) and found that it was able to detect shifts in a process mean more quickly.

The ARL is a popular measure for comparing control chart performance. It is the average number of observations until the first observation is detected outside the control limits. There are two components: $\mathrm{ARL}_{0}$ is called an incontrol ARL and ARL $_{1}$ is called an out-of-control ARL. $\mathrm{ARL}_{0}$ is the average number of observations before an out-of-control observation is detected when the process is incontrol and should be large while $\mathrm{ARL}_{1}$ is the average number of observations before an out-of-control signal is received when the process has shifted to the out-of-control state and should be small [9]. Several methods to calculate the ARL for many control charts, such as explicit formulas, Monte Carlo simulation, Markov chain, Martingale, and numerical integration equations (NIEs) methods [10]. Crowder [11] used a Fredholm integral equation to develop an approximation for the ARL of a Gaussian process on an EWMA control chart. Champ and Rigdon [12] employed the NIE and Markov chain approaches for the ARL of processes on CUSUM and EWMA control charts. Various researchers have aimed at approximating the ARL to measure the performance of control charts by using different methods. Robert [6] introduced the standard EWMA control chart using Monte Carlo simulation to evaluate the ARL. Areepong and Novikov [13] presented an explicit formula for the ARL and the average delay for a process running on an EWMA control chart while assuming that the observations follow an exponential distribution by using the Martingale approach. Phanyaem et al. [14] used a Fredholm integral equation technique to derive an exact expression of the ARL for the first-order autoregressive moving average (ARMA(1,1)) process running on the CUSUM control chart and compare the performance of control charts with the exact expression for the EWMA control chart. Sukparungsee and Areepong [15] derived explicit formulas for the ARL on an EWMA control chart for an autoregressive of order p (AR(p)) process. Sunthornwat et
al. [16] solved explicit formulas and optimal parameters for evaluating the ARL on an EWMA control chart for a longmemory AR fractionally integrated moving average (MA) (ARFIMA) process. Peerajit and Areepong [17] derived an exact solution for the ARL for an ARMA process with exogenous variables (ARMAX(p.q.r)) with exponential white noise running on a CUSUM control chart. Supharakonsakun et al. [18] suggested explicit formulas for the ARL of an MA(1) process running on a modified EWMA control chart. Phanthuna et al. [19] studied the run length distribution for the ARL of a stationary $\operatorname{AR}(\mathrm{p})$ process running on a modified EWMA control chart. After that, Silpakob et al. [20] derived an exact solution for the ARL of AR with explanatory variables (ARX(p,r)) processes running on a modified EWMA control chart. Most recently, Phanthuna and Areepong [21] studied the detection sensitivity of a modified EWMA control chart with a time series model for integrated MA (IMA) and fractional integrated MA (FIMA) models.

The main purpose of the present study is to derive explicit formulas for the ARL of an ARMA process with explanatory variables (ARMAX(p,q,r)) with exponential white noise running on a modified EWMA control chart based on Khan et al.'s [7] derivation. We apply Fredholmtype integral equations to derive an exact equation for two components of the ARL. This paper is organized as follows. An introduction to the control charts is provided in Section II. The explicit formulas and the NIE for the ARL of the process on the modified EWMA control chart are shown in Sections III and IV. Next, numerical results for comparing the performances of the ARLs derived by using integral equations and the NIE method are offered in Sections $V$ and VI, respectively. The practical application of the presented explicit formulas with real data is reported in Section VII. Finally, conclusions are given in Section VIII.

## II. Properties of the Control Charts Used in the Study

## A. The CUSUM Control Chart

This has been widely used to detect small shifts in process means in the same way as the EWMA control chart [5]. The CUSUM control chart can be defined as
$C_{t}=\max \left\{0, C_{t-1}+Y_{t}-a\right\} \quad ; t=1,2,3, \ldots$,
where $C_{t}$ is the CUSUM statistic, $Y_{t}$ is the sequence of the $\operatorname{ARMAX}(\mathrm{p}, \mathrm{q}, \mathrm{r})$ process with exponential white noise, $a$ is a constant. $C_{0}=u$ is the initial value when $u \in[0, b]$, where 0 is the lower control limit (LCL) and $b$ is the upper control limit (UCL).

## B. The Standard and Modified EWMA Control Charts

The modified EWMA control chart by defined as [7]
$M_{t}=(1-\lambda) M_{t-1}+\lambda Y_{t}+k\left(Y_{t}-Y_{t-1}\right) \quad ; t=1,2,3, \ldots$,
where $M_{t}$ is the modified EWMA statistic, $Y_{t}$ is the sequence of the $\operatorname{ARMAX}(\mathrm{p}, \mathrm{q}, \mathrm{r})$ process with exponential white noise, $\lambda$ is an exponential smoothing parameter
$(0<\lambda \leq 1)$, and $k$ is a constant $(k>0)$. Meanwhile, mean $E\left(M_{t}\right)=\mu_{0} \quad$ and variance $\quad \operatorname{Var}\left(M_{t}\right)=\frac{\left(\lambda+2 \lambda k+2 k^{2}\right)}{(2-\lambda)} \sigma^{2}$. From (2), the modified EWMA statistic is reduced to the standard EWMA statistic in [6] when $k=0$ (i.e., $\left.M_{t}=(1-\lambda) M_{t-1}+\lambda Y_{t}\right)$ and reduced to the primary modified EWMA statistic in [8] when $k=1$ (i.e., $\left.M_{t}=(1-\lambda) M_{t-1}+\lambda Y_{t}+\left(Y_{t}-Y_{t-1}\right)\right)$. Thus, we can derive the LCL and UCL of the two EWMA control charts as follows.

The respective LCL and UCL of the standard EWMA control chart with a control width limit $L_{S}$ are

$$
\begin{align*}
\mathrm{LCL} & =\mu_{0}-L_{S} \sigma \sqrt{\frac{\lambda}{2-\lambda}}  \tag{3a}\\
\text { and } \mathrm{UCL} & =\mu_{0}+L_{S} \sigma \sqrt{\frac{\lambda}{2-\lambda}},
\end{align*}
$$

while the respective LCL and UCL of the modified EWMA control chart with a control width limit $L_{M}$ are

$$
\begin{align*}
\mathrm{LCL} & =\mu_{0}-L_{M} \sigma \sqrt{\frac{\left(\lambda+2 \lambda k+2 k^{2}\right)}{(2-\lambda)}}  \tag{4a}\\
\text { and UCL } & =\mu_{0}+L_{M} \sigma \sqrt{\frac{\left(\lambda+2 \lambda k+2 k^{2}\right)}{(2-\lambda)}}, \tag{4b}
\end{align*}
$$

where $\mu_{0}$ is the target mean, $\sigma$ is the standard deviation of process, and $L_{S}, L_{M}>0$.

## III. Explicit Formulas for the ARL of the Process

## A. The ARMAX $(p, q, r)$ Process

This is defined as

$$
\begin{align*}
Y_{t}= & \omega+\phi_{1} Y_{t-1}+\phi_{2} Y_{t-2}+\ldots+\phi_{p} Y_{t-p}+\varepsilon_{t}-\theta_{1} \varepsilon_{t-1} \\
& -\theta_{2} \varepsilon_{t-2}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l} \quad ; t=1,2,3, \ldots, \tag{5}
\end{align*}
$$

where $\omega$ is a constant $(\omega \geq 0), \phi_{i}$ is an AR coefficient for $i=1,2, \ldots, p\left(\left|\phi_{i}\right|<1\right), \quad \theta_{j}$ is a MA coefficient for $j=1,2, \ldots, q\left(\left|\theta_{j}\right|<1\right), \varepsilon_{t}$ are independent and identically distributed (iid) observations in an exponential distribution $\left(\varepsilon_{t} \sim \operatorname{Exp}(\alpha)\right), X_{t l}$ are explanatory variables of $Y_{t}$, and $B_{l}$ is a coefficient for $l=1,2, \ldots, r$. The initial value for the ARMAX (p,q,r) process is 1 .

## B. Explicit Formulas

Explicit formulas for the ARL of an ARMAX(p,q,r) process running on the CUSUM control chart are shown in [17]. Explicit formulas for the ARL of an ARMAX(p,q,r) process on the modified EWMA control chart are derived as follows:

$$
\begin{aligned}
M_{t}= & (1-\lambda) M_{t-1}+(\lambda+k) \varepsilon_{t}-k Y_{t-1} \\
& +(\lambda+k)\binom{\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}}{-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l}}
\end{aligned}
$$

If $Y_{1}$ signals the out-of-control state for $M_{1}$ when $M_{0}=u$, then

$$
\begin{aligned}
M_{1}= & (1-\lambda) u+(\lambda+k) \varepsilon_{1}-k Y_{0} \\
& +(\lambda+k)\binom{\omega+\phi_{1} Y_{0}+\ldots+\phi_{p} Y_{t-p}}{-\theta_{1} \varepsilon_{0}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l}}
\end{aligned}
$$

If $\varepsilon_{1}$ is the in-control state for $M_{1}$, then $0 \leq M_{1} \leq b ; 0$ is LCL and $b$ is UCL. Consider the Fredholm integral equation of the second kind [22] following

$$
\begin{equation*}
H(u)=1+\int H\left(M_{1}\right) f\left(\varepsilon_{1}\right) d\left(\varepsilon_{1}\right) \tag{6}
\end{equation*}
$$

Moreover, $H(u)$ can be written as

$$
\left.H(u)=1+\int_{0}^{b} L\left\{\begin{array}{l}
(1-\lambda) u-k Y_{t-1}+(\lambda+k) y \\
+(\lambda+k)\binom{\omega+\phi_{1} Y_{t-1}+\ldots}{-\theta_{1} \varepsilon_{t-1}-\ldots+\sum_{l=1}^{r} \beta_{l} X_{t l}}
\end{array}\right)\right\} f(y) d y .
$$

Let

$$
\begin{aligned}
w= & (1-\lambda) u-k Y_{t-1}+(\lambda+k) y \\
& +(\lambda+k)\binom{\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}}{-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l}}
\end{aligned}
$$

By changing the variable, we can obtain the integral equation as follows:

$$
\begin{align*}
H(u)= & 1+\frac{1}{\lambda+k} \\
& \int_{0}^{b} H(w) f\left\{\begin{array}{l}
\frac{w-(1-\lambda) u}{\lambda+k}+\frac{k Y_{t-1}}{\lambda+k} \\
-\binom{\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}}{-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l}}
\end{array}\right) \tag{7}
\end{align*}
$$

If $Y_{t} \sim \operatorname{Exp}(\alpha)$ and $f(y)=\frac{1}{\alpha} e^{\frac{-y}{\alpha}} ; y \geq 0$, then

$$
\begin{align*}
H(u)= & 1+\frac{1}{\lambda+k} \cdot \\
& \int_{0}^{b} H(w) \frac{1}{\alpha} e^{-\frac{1}{\alpha}\left\{\frac{w-(1-\lambda) u}{\lambda+k}+\frac{k Y_{t-1}}{\lambda+k}-\left(-\theta_{1} \varepsilon_{t-1} \cdots+\theta_{q} \varepsilon_{t-q}+\sum_{l=1} \beta_{l} x_{t l}\right)\right\}} d w \tag{8}
\end{align*}
$$

 we obtain
$H(u)=1+\frac{F(u)}{\alpha(\lambda+k)} \int_{0}^{b} H(w) e^{\frac{-w}{\alpha(\lambda+k)}} d w \quad ; 0 \leq u \leq b$.

Let $B=\int_{0}^{b} H(w) e^{\frac{-w}{\alpha(\lambda+k)}} d w$, then $H(u)=1+\frac{F(u)}{\alpha(\lambda+k)} \cdot B$.
Consequently, we obtain

$$
\begin{equation*}
H(u)=1+\frac{1}{\alpha(\lambda+k)} e^{\frac{(1-\lambda) u-k Y_{t-1}+\frac{1}{\alpha}}{\alpha(\lambda+k)}\binom{\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}}{-\theta_{1} \varepsilon_{t-1} \cdots \theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{l l}}} \cdot B \tag{9}
\end{equation*}
$$

By solving for constant $B$, we obtain

$$
\begin{aligned}
& B=\int_{0}^{b} H(w) e^{\frac{-w}{\alpha(\lambda+k)}} d w \\
&=\int_{0}^{b}\left[1+\frac{B}{\alpha(\lambda+k)} F(w)\right] e^{\frac{-w}{\alpha(\lambda+k)}} d w \\
&=\int_{0}^{b} e^{\frac{-w}{\alpha(\lambda+k)}} d w+\int_{0}^{b} \frac{\left.B e^{\frac{(1-\lambda) w-k Y_{t-1}}{\alpha(\lambda+k)}+\frac{1}{\alpha}\left(-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1} \beta_{1} x_{l l}\right.}\right)}{\alpha(\lambda+k)} \cdot e^{\frac{-w}{\alpha(\lambda+k)}} d w \\
&=\frac{-\alpha(\lambda+k)\left(e^{\frac{-b}{\alpha(\lambda+k)}}-1\right)}{\lambda} \\
&\left.1+\frac{\left.e^{\frac{-k Y_{t-1}}{\alpha(\lambda+k)}+\frac{1}{\alpha}\left(\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} x_{l l}\right.}\right)}{\lambda} e^{\frac{-\lambda b}{\alpha(\lambda+k)}}-1\right)
\end{aligned}
$$

By substituting constant $B$ into Eq. (23), we arrive at

$$
\begin{align*}
& H(u)=1+\frac{e^{\frac{(1-\lambda) u-k Y_{t-1}}{\alpha(\lambda+k)}+\frac{1}{\alpha}\left(\omega+\phi_{1} Y_{t-1}+\ldots+\phi_{p} Y_{t-p}-\theta_{1} \varepsilon_{t-1}-\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{l l}\right)}}{\alpha(\lambda+k)} \\
&\binom{-\lambda \alpha(\lambda+k)\left[e^{\frac{-b}{\alpha(\lambda+k)}}-1\right]}{\left.\lambda+e^{\frac{-k Y_{t-1}+\frac{1}{\alpha(\lambda+k)} \alpha\left(-\theta_{1} \varepsilon_{t-1}+\ldots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{l} X_{t l}\right)}{\omega+Y_{t-1}+\ldots+\phi_{t-p} Y^{\frac{-\lambda b}{\alpha(\lambda+k)}}-1} e^{\alpha}}\right)} . \tag{10}
\end{align*}
$$

Hence, the one-sided explicit formulas for the ARL on a modified EWMA control chart for an ARMAX(p,q,r) process can be derived by using the Fredholm integral equation of the second kind. Let $\alpha=\alpha_{0}$ for the process is in the in-control state, and $\alpha=\alpha_{1}$ for the process is in the out-of-control state, the one-sided explicit formulas for $A R L_{0}$ and $A R L_{1}$ can be written as follows:

and


## C. The Existence and Uniqueness of the Explicit Formulas

Here, we show the existence and uniqueness of the solution in (8). First, we define

Theorem 1. Banach's fixed-point theorem [23].
Let $C[0, b]$ be a set of all of the continuous functions on complete metric $(X, d)$, and assume that $T: X \rightarrow X$ is a contraction mapping with contraction constant $0 \leq s<1$; i.e., $\quad\left\|T\left(H_{1}\right)-T\left(H_{2}\right)\right\| \leq s\left\|H_{1}-H_{2}\right\| \forall H_{1}, H_{2} \in X$.

Subsequently, $H(.) \in X$ is unique at $T(H(u))=H(u)$; i.e., it has a unique fixed point in $X$.

Proof: To show that $T$ defined in (13) is a contraction mapping for $H_{1}, H_{2} \in C[0, b]$, we use the inequality $\left\|T\left(H_{1}\right)-T\left(H_{2}\right)\right\| \leq s\left\|H_{1}-H_{2}\right\| \forall H_{1}, H_{2} \in C(0, b)$ with $0 \leq s<1$. Consider (8) and (13), then

$$
\begin{aligned}
& \left\|T\left(H_{1}\right)-T\left(H_{2}\right)\right\|_{\infty}=\sup _{u \in[0, b]}\left|\frac{F(u)}{\alpha(\lambda+k)} \int_{0}^{b}\left(H_{1}(w)-H_{2}(w)\right) e^{\frac{-w}{\alpha(\lambda+k)}} d w\right| \\
& \quad \leq \sup _{u \in[0, b]}\left|\left\|H_{1}-H_{2}\right\|_{\infty} F(u)\left(1-e^{\frac{-b}{\alpha(\lambda+k)}}\right)\right| \\
& \quad=\left\|L_{1}-L_{2}\right\|_{\infty}\left|1-e^{\frac{-b}{\alpha(\lambda+k)}}\right| \sup _{u \in[0, b]}|F(u)| \\
& \quad \leq s\left\|L_{1}-L_{2}\right\|_{\infty},
\end{aligned}
$$

where $s=\left|1-e^{\frac{-b}{\alpha(\lambda+k)}}\right| \sup _{u \in[0, b]}|F(u)|$ and
$F(u)=e^{\frac{(1-\lambda) u-k Y_{t-1}+\frac{1}{\alpha}\left(\omega+\phi Y_{1-1}+\ldots+\phi_{p} Y_{1-p}-\theta \theta_{1} \varepsilon_{t-1} \cdots-\theta_{q} \varepsilon_{t-q}+\sum_{l=1}^{r} \beta_{1} X_{U}\right)}{\alpha(\lambda+k)}} ; 0 \leq s<1$.
Therefore, as confirmed by applying Banach's fixed-point theorem, the solution exists and is unique.

## IV. The NIE for the ARL of the Process

The NIE approach is widely used for evaluating the ARL. It can be based on one of various quadrature rules (midpoint, trapezoidal, Simpson's, and Gauss-Legendre), all of which give ARLs that are very close to each other [24]. In the present study, we use the Gauss-Legendre rule to evaluate the ARL. The Fredholm integral equation of the second kind for the ARL for the ARMAX(p,q,r) process running on the modified EWMA control chart in (10) can be evaluated using the quadrature formula. We apply the Gauss-Legendre rule as follows:

Given that

$$
f\left(a_{j}\right)=f\left\{\begin{array}{l}
\frac{a_{j}-(1-\lambda) a_{i}}{(\lambda+k)}+\frac{k Y_{t-1}}{(\lambda+k)}-\omega-\phi_{1} Y_{t-1}-\ldots  \tag{14}\\
-\phi_{p} Y_{t-p}+\theta_{1} \varepsilon_{t-1}+\ldots+\theta_{q} \varepsilon_{t-q}-\sum_{l=1}^{r} \beta_{l} X_{t l}
\end{array}\right\},
$$

the estimation for the integral equation by using GaussLegendre rule is in the form

$$
\begin{equation*}
\int_{0}^{b} H(w) f(w) d w \approx \sum_{j=1}^{m} w_{j} f\left(a_{j}\right), \tag{15}
\end{equation*}
$$

where $a_{j}=\frac{b}{m}\left(j-\frac{1}{2}\right)$ and $w_{j}=\frac{b}{m} ; j=1,2, \ldots, m$.
The numerical approximation $\tilde{H}(u)$ for the integral equations can be found as the solution to the following equations:

$$
\begin{aligned}
\tilde{H}\left(a_{i}\right)= & 1+\frac{1}{\lambda+k} \sum_{j=1}^{m} w_{j} \tilde{H}\left(a_{j}\right) \cdot \\
& f\left\{\begin{array}{l}
\frac{a_{j}-(1-\lambda) a_{i}}{(\lambda+k)}+\frac{k Y_{t-1}}{(\lambda+k)}-\omega-\phi_{1} Y_{t-1}-\ldots \\
-\phi_{p} Y_{t-p}+\theta_{1} \varepsilon_{t-1}+\ldots+\theta_{q} \varepsilon_{t-q}-\sum_{l=1}^{r} \beta_{l} X_{t l}
\end{array}\right\} \\
\tilde{H}\left(a_{m}\right)= & 1+\frac{1}{\lambda+k} \sum_{j=1}^{m} w_{j} \tilde{H}\left(a_{j}\right) \cdot \\
& f\left\{\begin{array}{l}
\frac{a_{j}-(1-\lambda) a_{m}}{(\lambda+k)}+\frac{k Y_{t-1}}{(\lambda+k)}-\omega-\phi_{1} Y_{t-1}-\ldots \\
-\phi_{p} Y_{t-p}+\theta_{1} \varepsilon_{t-1}+\ldots+\theta_{q} \varepsilon_{t-q}-\sum_{l=1}^{r} \beta_{l} X_{t l}
\end{array}\right\}
\end{aligned}
$$

This set of $m$ equations with $m$ unknowns can be rewritten in matrix form. The column vector of $\tilde{H}\left(a_{i}\right)$ is $\mathbf{L}_{m \times 1}=\left(\tilde{H}\left(a_{1}\right), \tilde{H}\left(a_{2}\right), \ldots, \tilde{H}\left(a_{m}\right)\right)^{\prime}$. Since $\mathbf{1}_{m \times 1}=(1,1, \ldots, 1)^{\prime}$ is a column vector of ones and $\mathbf{R}_{m \times m}$ is a matrix, we can define $m$ to $m^{\text {th }}$ composition of matrix $\mathbf{R}$ as follows:
$\left[R_{i j}\right] \approx \frac{1}{\lambda+k} w_{j} f\left\{\begin{array}{l}\frac{a_{j}-(1-\lambda) a_{i}}{(\lambda+k)}+\frac{k Y_{t-1}}{(\lambda+k)}-\omega-\phi_{1} Y_{t-1}-\ldots \\ -\phi_{p} Y_{t-p}+\theta_{1} \varepsilon_{t-1}+\ldots+\theta_{q} \varepsilon_{t-q}-\sum_{l=1}^{r} \beta_{l} X_{t l}\end{array}\right\}$,
and $\quad \mathbf{I}_{m}=\operatorname{diag}(1,1, \ldots, 1)$ as a unit matrix order $m$. If $(\mathbf{I}-\mathbf{R})^{-1}$ exists, the numerical approximation for the integral equation in matrix terms can be written as $\mathbf{G}_{m \times 1}=\left(\mathbf{I}_{m}-\mathbf{R}_{m \times m}\right)^{-1} \mathbf{1}_{m \times 1}$.
Finally, by substituting $a_{i}$ with $u$ in $\tilde{H}\left(a_{i}\right)$, the numerical integration equation for function $\tilde{H}(u)$ can be obtained as

$$
\begin{align*}
\tilde{H}(u)= & 1+\frac{1}{\lambda+k} \sum_{j=1}^{m} w_{j} \tilde{H}\left(a_{j}\right) \cdot \\
& f\left\{\begin{array}{l}
\frac{a_{j}-(1-\lambda) u}{(\lambda+k)}+\frac{k Y_{t-1}}{(\lambda+k)}-\omega-\phi_{1} Y_{t-1}-\ldots \\
-\phi_{p} Y_{t-p}+\theta_{1} \varepsilon_{t-1}+\ldots+\theta_{q} \varepsilon_{t-q}-\sum_{l=1}^{r} \beta_{l} X_{t l}
\end{array}\right\} . \tag{16}
\end{align*}
$$

## V. Comparison of the Efficacies of the NIE and Explicit Formulas Methods

Here, the details of the simulation study to compare the efficacies for the ARL on the modified EWMA control chart of an ARMAX(p,q,r) process for the explicit formulas $(H(u))$ and the NIE method $(\tilde{H}(u))$ are provided. The parameter of the modified EWMA control chart ( $\lambda=0.05$, 0.1 and $k=1$ ) and $\operatorname{ARMAX}(\mathrm{p}, \mathrm{q}, \mathrm{r})$ process with the incontrol process $\alpha_{0}=1$; where the shift size ( $\delta$ ) varied as $0.001,0.003,0.005,0.007,0.01,0.03,0.05,0.07,0.1,0.3$, or 0.5 . given $A R L_{0}=370$. The absolute percentage difference between the ARL methods is defined as

$$
\begin{equation*}
\operatorname{Diff}(\%)=\frac{|H(u)-\tilde{H}(u)|}{H(u)} \times 100 . \tag{17}
\end{equation*}
$$

Equations (10) and (16) were used to evaluate the ARL on the modified EWMA control chart for the TABLE I
Comparison of the ARL for an Armax $(1,1,1)$ Process on the Modified EWMA Control Chart by Using Explicit Formulas and the Nie METHOD wITH $\hat{\omega}=2$.

| $\lambda$ | $\delta$ | $\hat{\phi}=0.1, \hat{\theta}=-0.1, \hat{\beta}=0.1$ and $b=0.546791$ |  |  |  | $\hat{\phi}=0.2, \hat{\theta}=-0.2, \hat{\beta}=0.2$ and $b=0.404322$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Explicit | NIE | Time ${ }^{\text {a }}$ | Diff\% | Explicit | NIE | Time | Diff\% |
| 0.05 | 0.00 | 370 | 370 | 10.812 | 0.00000000 | 370 | 370 | 10.531 | 0.00000000 |
|  | 0.001 | 229.904260 | 229.904259 | 11.297 | 0.00000041 | 221.949849 | 221.949848 | 10.531 | 0.00000022 |
|  | 0.003 | 130.988167 | 130.988167 | 12.750 | 0.00000036 | 123.419607 | 123.419607 | 10.751 | 0.00000019 |
|  | 0.005 | 91.687880 | 91.687879 | 11.484 | 0.00000034 | 85.574008 | 85.574008 | 11.531 | 0.00000018 |
|  | 0.007 | 70.589029 | 70.589028 | 11.625 | 0.00000032 | 65.549877 | 65.549876 | 11.468 | 0.00000017 |
|  | 0.01 | 52.539577 | 52.539577 | 10.979 | 0.00000031 | 48.579748 | 48.579748 | 11.063 | 0.00000017 |
|  | 0.03 | 19.696613 | 19.696612 | 10.750 | 0.00000027 | 18.073970 | 18.073970 | 10.687 | 0.00000015 |
|  | 0.05 | 12.302525 | 12.302525 | 11.328 | 0.00000025 | 11.273029 | 11.273029 | 10.609 | 0.00000014 |
|  | 0.07 | 9.042091 | 9.042091 | 10.593 | 0.00000023 | 8.282850 | 8.282850 | 10.469 | 0.00000013 |
|  | 0.10 | 6.563506 | 6.563506 | 12.156 | 0.00000021 | 6.014209 | 6.014209 | 10.467 | 0.00000011 |
|  | 0.30 | 2.688419 | 2.688419 | 11.328 | 0.00000011 | 2.483831 | 2.483831 | 11.687 | 0.00000006 |
|  | 0.50 | 1.937998 | 1.937998 | 10.656 | 0.00000007 | 1.808533 | 1.808533 | 11.656 | 0.00000003 |
| 0.1 |  |  | $b=0.550849$ |  |  |  | $b=0$ | 6401 |  |
|  | 0.00 | 370 | 370 | 10.438 | 0.00000000 | 370 | 370 | 11.188 | 0.00000000 |
|  | 0.001 | 227.026409 | 227.026407 | 11.344 | 0.00000076 | 218.788511 | 218.788510 | 11.078 | 0.00000040 |
|  | 0.003 | 128.207062 | 128.207061 | 11.531 | 0.00000055 | 120.550648 | 120.550647 | 11.031 | 0.00000028 |
|  | 0.005 | 89.435741 | 89.435741 | 10.437 | 0.00000046 | 83.301325 | 83.301325 | 10.828 | 0.00000024 |
|  | 0.007 | 68.734285 | 68.734285 | 10.407 | 0.00000042 | 63.699735 | 63.699735 | 10.687 | 0.00000022 |
|  | 0.01 | 51.087218 | 51.087218 | 10.703 | 0.00000038 | 47.145277 | 47.145277 | 11.156 | 0.00000020 |
|  | 0.03 | 19.122925 | 19.122925 | 11.188 | 0.00000029 | 17.518459 | 17.518459 | 10.969 | 0.00000016 |
|  | 0.05 | 11.952362 | 11.952362 | 11.281 | 0.00000026 | 10.936191 | 10.936191 | 11.531 | 0.00000014 |
|  | 0.07 | 8.793327 | 8.793327 | 11.156 | 0.00000024 | 8.044608 | 8.044608 | 11.313 | 0.00000013 |
|  | 0.10 | 6.392839 | 6.392839 | 10.860 | 0.00000021 | 5.851618 | 5.851618 | 11.484 | 0.00000011 |
|  | 0.30 | 2.640331 | 2.640331 | 11.562 | 0.00000011 | 2.439131 | 2.439131 | 11.078 | 0.00000005 |
|  | 0.50 | 1.912903 | 1.912903 | 10.906 | 0.00000006 | 1.785612 | 1.785612 | 11.047 | 0.00000003 |

${ }^{\text {a }}$ The calculations for the NIE method are based on Windows 10 Professional with an Intel Core i5 CPU with number of nodes 1000 iterations TABLE II
Comparison of the ArL for an ARMAX $(1,2,2)$ Process on the Modified EWMA Control Chart by Using Explicit Formulas and the nie METHOD WITH $\hat{\omega}=2, \hat{\phi}=0.1$.

| $\lambda$ | $\delta$ | $\hat{\theta}_{1}=-0.1, \hat{\theta}_{2}=0.2, \hat{\beta}_{1}=0.1, \hat{\beta}_{2}=-0.1$ and $b=0.739943$. |  |  |  | $\hat{\theta}_{1}=-0.3, \hat{\theta}_{2}=0.2, \hat{\beta}_{1}=0.1, \hat{\beta}_{2}=-0.2$ and $b=0.6689122$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Explicit | NIE | Time | Diff\% | Explicit | NIE | Time | Diff\% |
| 0.05 | 0.00 | 370 | 370 | 10.937 | 0.00000000 | 370.000000 | 370.000000 | 11.515 | 0.00000000 |
|  | 0.001 | 238.302931 | 238.302929 | 11.672 | 0.00000078 | 235.456965 | 235.456964 | 12.063 | $0.00000063$ |
|  | 0.003 | 139.365880 | 139.365879 | 11.703 | 0.00000067 | 136.474732 | 136.474731 | 11.828 | 0.00000055 |
|  | 0.005 | 98.588803 | 98.588802 | 12.156 | 0.00000063 | 96.190221 | 96.190221 | 11.875 | 0.00000051 |
|  | 0.007 | 76.337605 | 76.337605 | 11.390 | 0.00000060 | 74.331849 | 74.331849 | 12.36 | 0.00000049 |
|  | 0.01 | 57.098633 | 57.098632 | 11.281 | 0.00000058 | 55.502639 | 55.502639 | 12.078 | 0.00000047 |
|  | 0.03 | 21.598196 | 21.598196 | 11.625 | 0.00000051 | 20.928321 | 20.928321 | 11.406 | 0.00000041 |
|  | 0.05 | 13.515006 | 13.515005 | 11.110 | 0.00000047 | 13.087170 | 13.087170 | 11.344 | 0.00000038 |
|  | 0.07 | 9.938871 | 9.938871 | 11.235 | 0.00000043 | 9.622141 | 9.622141 | 12.016 | 0.00000035 |
|  | 0.10 | 7.214336 | 7.214336 | 11.678 | 0.00000039 | 6.984252 | 6.984252 | 11.532 | 0.00000032 |
|  | 0.30 | 2.934234 | 2.934234 | 11.969 | 0.00000021 | 2.846976 | 2.846976 | 12.344 | 0.00000017 |
|  | 0.50 | 2.095283 | 2.095283 | 11.781 | 0.00000012 | 2.039270 | 2.039270 | 11.813 | 0.00000010 |
| 0.1 |  |  | $b=0.747141$ |  |  |  | $b=0$ | 48496 |  |
|  | 0.00 | 370 | $370$ | 11.765 | 0.00000000 | 370.000000 | 370.000000 | 12.031 | 0.00000000 |
|  | $0.001$ | 235.720048 | 235.720045 | 11.594 | 0.00000146 | 232.736927 | 232.736925 | 11.875 | 0.00000117 |
|  | 0.003 | 136.733055 | 136.733053 | 11.453 | 0.00000106 | 133.772024 | 133.772023 | 11.328 | 0.00000085 |
|  | 0.005 | 96.410183 | 96.410182 | 11.891 | 0.00000089 | 93.973650 | 93.973649 | 10.906 | 0.00000072 |
|  | 0.007 | 74.522127 | 74.522127 | 11.906 | 0.00000080 | 72.493444 | 72.493443 | 11.297 | 0.00000064 |
|  | 0.01 | 55.662295 | 55.662295 | 11.891 | 0.00000072 | 54.054049 | 54.054049 | 11.703 | 0.00000058 |
|  | 0.03 | 21.018871 | 21.018871 | 11.469 | 0.00000055 | 20.348695 | 20.348695 | 11.563 | 0.00000044 |
|  | 0.05 | 13.158986 | 13.158986 | 11.328 | 0.00000048 | 12.731843 | 12.731843 | 12.218 | 0.00000039 |
|  | 0.07 | 9.684822 | 9.684822 | 11.593 | 0.00000044 | 9.368979 | 9.368979 | 11.235 | 0.00000036 |
|  | 0.10 | 7.039143 | 7.039143 | 11.344 | 0.00000039 | 6.809976 | 6.809976 | 11.547 | 0.00000032 |
|  | 0.30 | 2.883711 | 2.883711 | 11.594 | 0.00000021 | 2.797098 | 2.797098 | 11.938 | 0.00000017 |
|  | 0.50 | 2.068491 | 2.068491 | 11.406 | 0.00000012 | 2.012954 | 2.012954 | 11.781 | 0.00000010 |

ARMAX( $\mathrm{p}, \mathrm{q}, \mathrm{r}$ ) process with exponential white noise. The results are shown in Tables I and II.
The results in Tables I and II shows that the ARLs derived by the explicit formulas are close to the NIE method (Diff(\%) was less than $0.00001 \%$ ). However, the computational time for the NIE method was more than 10 seconds while that of the explicit formulas was less than 1 second.

## VI. The ARL Results Using Various Control Charts

The performances of the ARL derived using explicit formulas for an ARMAX(p,q,r) process with exponential white noise running on the standard and modified EWMA and CUSUM control charts were compared by using the relative mean index (RMI). The ARL with the lowest value indicates the best performance. The RMI is defined as
$R M I(r)=\frac{1}{n} \sum_{i=1}^{n}\left(\frac{A R L_{i}(r)-\operatorname{Min}\left[A R L_{i}(s)\right]}{\operatorname{Min}\left[A R L_{i}(s)\right]}\right)$,
where $A R L_{i}(r)$ is the ARL of the control chart for the shift size in row $i$ and $\operatorname{Min}\left[A R L_{i}(s)\right]$ denotes the smallest ARL of the three control charts in comparison to the shift size in row $i$, for $i=1,2, \ldots, n$. The control chart with the smallest RMI is the best for a particular set of criteria.

For the simulation study with $\operatorname{ARMAX}(1,1,1)$ and ARMAX $(2,2,2)$ processes, the parameter of the modified EWMA control chart ( $\lambda=0.05,0.1$ and $k=1,2,3$ ) and the in-control process $\alpha_{0}=1$; where $\delta=0.001,0.003,0.005$, $0.007,0.01,0.03,0.05,0.07,0.1,0.3$, or 0.5 and $A R L_{0}=370$. The results for the $\operatorname{ARMAX}(1,1,1)$ processes are reported in Table III and plot in Fig. 1 while those for the $\operatorname{ARMAX}(2,2,2)$ process are reported in Table IV and plot in Fig. 2.

TABLE III
COMPARISON OF THE ARL FOR THE ARMAX $(1,1,1)$ Process on CUSUM, STANDARd, AND MODIFIED EWMA Control Charts with $\hat{\omega}=2, \hat{\phi}=0.2$, AND $\hat{\beta}=0.1$.

| $\lambda$ | $\hat{\theta}$ | $\delta$ | CUSUM | EWMA | Modified EWMA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | $k=1$ | $k=2$ | $k=3$ |
| 0.05 | 0.2 |  | $a=5, b=3.1466$ | $b=0.00000001266$ | $b=0.3339873$ | $b=0.6689124$ | $b=1.003798$ |
|  |  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  |  | 0.001 | 367.719 | 361.985 | 269.720 | 235.469 | 222.534 |
|  |  | 0.003 | 363.149 | 346.437 | 174.876 | 136.479 | 124.045 |
|  |  | 0.005 | 358.654 | 331.615 | 129.375 | 96.192 | 86.130 |
|  |  | 0.007 | 354.231 | 317.482 | 102.661 | 74.333 | 66.050 |
|  |  | 0.01 | 347.731 | 297.502 | 78.381 | 55.503 | 49.021 |
|  |  | 0.03 | 308.174 | 194.794 | 30.426 | 20.928 | 18.381 |
|  |  | 0.05 | 274.357 | 129.684 | 18.898 | 13.087 | 11.542 |
|  |  | 0.07 | 245.299 | 87.737 | 13.728 | 9.622 | 8.531 |
|  |  | 0.10 | 208.949 | 50.275 | 9.766 | 6.984 | 6.242 |
|  |  | 0.30 | 86.574 | 3.180 | 3.554 | 2.847 | 2.651 |
|  |  | 0.50 | 45.429 | 1.217 | 2.373 | 2.039 | 1.944 |
|  | RMI |  | 15.246 | 4.210 | 0.503 | 0.150 | 0.050 |
|  | -0.2 |  | $a=5, b=3.6681$ | $b=0.000000008486$ | $b=0.2232272$ | $b=0.447094$ | $b=0.6709292$ |
|  |  | 0.00 | $370$ | $370$ | 370 | 370 | 370 |
|  |  | 0.001 | 367.627 | 361.830 | 260.325 | 224.551 | 211.263 |
|  |  | 0.003 | 362.880 | 346.014 | 163.397 | 125.860 | 113.915 |
|  |  | 0.005 | 358.213 | 330.947 | 119.051 | 87.534 | 78.117 |
|  |  | 0.007 | 353.624 | 316.594 | 93.629 | 67.160 | 59.516 |
|  |  | 0.01 | 346.883 | 296.321 | 70.907 | 49.841 | 43.932 |
|  |  | 0.03 | 305.978 | 192.543 | 27.069 | 18.588 | 16.324 |
|  |  | 0.05 | 271.182 | 127.253 | 16.735 | 11.599 | 10.237 |
|  |  | 0.07 | 241.426 | 85.494 | 12.128 | 8.523 | 7.567 |
|  |  | 0.10 | 204.414 | 48.514 | 8.611 | 6.188 | 5.542 |
|  |  | 0.30 | 82.204 | 2.988 | 3.143 | 2.548 | 2.384 |
|  |  | 0.50 | 42.452 | 1.190 | 2.123 | 1.849 | 1.771 |
|  | RMI |  | 16.378 | 4.715 | 0.494 | 0.141 | 0.041 |
| 0.1 | 0.2 |  | $a=5, b=3.1466$ | $b=0.00053475$ | $b=0.3376842$ | $b=0.6748497$ | $b=1.012368$ |
|  |  |  | 370 | 370 | 370 | 370 | 370 |
|  |  | 0.001 | 367.719 | 365.532 | 263.023 | 232.743 | 221.178 |
|  |  | 0.003 | 363.149 | 356.751 | 166.655 | 133.774 | 122.783 |
|  |  | 0.005 | 358.654 | 348.215 | 121.983 | 93.975 | 85.123 |
|  |  | 0.007 | 354.231 | 339.915 | 96.207 | 72.494 | 65.227 |
|  |  | 0.01 | 347.731 | 327.893 | 73.063 | 54.054 | 48.381 |
|  |  | 0.03 | 308.174 | 259.296 | 28.126 | 20.349 | 18.130 |
|  |  | 0.05 | 274.357 | 206.884 | 17.473 | 12.732 | 11.388 |
|  |  | 0.07 | 245.299 | 166.465 | 12.713 | 9.369 | 8.421 |
|  |  | 0.10 | 208.949 | 121.951 | 9.071 | 6.810 | 6.167 |
|  |  | 0.30 | 86.574 | 22.409 | 3.367 | 2.797 | 2.629 |
|  |  | 0.50 | 45.429 | 6.907 | 2.280 | 2.013 | 1.932 |
|  | RMI |  | 14.249 | 7.812 | 0.374 | 0.086 | 0.000 |
|  | -0.2 |  | $a=5, b=3.6681$ | $b=0.0003582$ | $b=0.2251526$ | $b=0.4498835$ | $b=0.674842$ |
|  |  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  |  | 0.001 | 367.627 | 365.432 | 253.108 | 221.486 | 209.568 |
|  |  | 0.003 | 362.880 | 356.357 | 155.096 | 123.016 | 112.452 |
|  |  | 0.005 | 358.213 | 347.543 | 111.810 | 85.264 | 76.981 |
|  |  | 0.007 | 353.624 | 338.980 | 87.418 | 65.305 | 58.601 |
|  |  | 0.01 | 346.883 | 326.589 | 65.872 | 48.398 | 43.229 |
|  |  | 0.03 | 305.978 | 256.213 | 24.965 | 18.026 | 16.055 |
|  |  | 0.05 | 271.182 | 202.873 | 15.446 | 11.257 | 10.073 |
|  |  | 0.07 | 241.426 | 162.053 | 11.216 | 8.281 | 7.451 |
|  |  | 0.10 | 204.414 | 117.499 | 7.992 | 6.022 | 5.463 |
|  |  | 0.30 | 82.204 | 20.484 | 2.982 | 2.502 | 2.361 |
|  |  | 0.50 | 42.452 | 6.160 | 2.044 | 1.825 | 1.759 |
|  | RMI |  | 15.620 | 8.608 | 0.378 | 0.087 | 0.000 |

IAENG International Journal of Applied Mathematics, 53:1, IJAM_53_1_32


Fig. 1. The ARL for an ARMAX ( $1,1,1$ ) process running on CUSUM, standard, and modified EWMA control charts. (a) The ARL for $\lambda=0.05$ and $\hat{\theta}=0.2$ and (b) the ARL for $\lambda=0.1$ and $\hat{\theta}=-0.2$.

TABLE IV
COMPARISON OF THE ARL FOR THE ARMAX $(2,2,2)$ PROCESS ON CUSUM, STANDARD, AND MODIFIED EWMA CONTROL CHARTS wITH $\hat{\omega}^{=}=2, \hat{\phi}_{1}=0.1$, $\hat{\phi}_{2}=0.2, \hat{\theta}_{2}=0.2, \hat{\beta}_{1}=0.1$ AND $\hat{\beta}_{2}=0.1$.

| $\lambda$ | $\hat{\theta}_{1}$ | $\delta$ | CUSUM | EWMA | Modified EWMA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | $k=1$ | $k=2$ | $k=3$ |
| 0.05 | 0.3 |  | $a=5, b=3.0274$ | $b=0.000000013991$ | $b=0.369456$ | $b=0.739944$ | $b=1.11039$ |
|  |  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  |  | 0.001 | 367.767 | 362.009 | 272.145 | 238.359 | 225.506 |
|  |  | 0.003 | 363.227 | 346.529 | 177.954 | 139.385 | 126.825 |
|  |  | 0.005 | 358.761 | 331.768 | 132.196 | 98.598 | 88.363 |
|  |  | 0.007 | 354.367 | 317.692 | 105.156 | 76.343 | 67.885 |
|  |  | 0.01 | 347.907 | 297.786 | 80.466 | 57.102 | 50.461 |
|  |  | 0.03 | 308.582 | 195.352 | 31.382 | 21.599 | 18.971 |
|  |  | 0.05 | 274.934 | 130.294 | 19.517 | 13.515 | 11.917 |
|  |  | 0.07 | 245.998 | 88.304 | 14.187 | 9.939 | 8.809 |
|  |  | 0.10 | 209.765 | 50.724 | 10.098 | 7.214 | 6.445 |
|  |  | 0.30 | 87.380 | 3.231 | 3.674 | 2.934 | 2.729 |
|  |  | 0.50 | 45.994 | 1.224 | 2.447 | 2.095 | 1.995 |
|  | RMI |  | 14.922 | 4.082 | 0.505 | 0.153 | 0.052 |
|  | -0.3 |  | $a=5, b=3.8159$ | $b=0.000000007679$ | $b=0.2018705$ | $b=0.4043218$ | $b=0.6067437$ |
|  |  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  |  | 0.001 | 367.599 | 361.820 | 258.046 | 221.933 | 208.579 |
|  |  | 0.003 | 362.787 | 345.935 | 160.712 | 123.414 | 111.594 |
|  |  | 0.005 | 358.056 | 330.807 | 116.679 | 85.572 | 76.309 |
|  |  | 0.007 | 353.405 | 316.397 | 91.576 | 65.548 | 58.054 |
|  |  | 0.01 | 346.575 | 296.050 | 69.225 | 48.579 | 42.801 |
|  |  | 0.03 | 305.176 | 192.000 | 26.329 | 18.074 | 15.872 |
|  |  | 0.05 | 270.024 | 126.662 | 16.261 | 11.273 | 9.951 |
|  |  | 0.07 | 240.019 | 84.949 | 11.778 | 8.283 | 7.356 |
|  |  | 0.10 | 202.776 | 48.088 | 8.360 | 6.014 | 5.390 |
|  |  | 0.30 | 80.684 | 2.943 | 3.054 | 2.484 | 2.326 |
|  |  | 0.50 | 41.449 | 1.184 | 2.069 | 1.809 | 1.734 |
|  | RMI |  | 16.600 | 4.841 | 0.492 | 0.139 | 0.039 |
| 0.1 | 0.3 |  | $a=5, b=3.0274$ | $b=0.0005912$ | $b=0.373837$ | $b=0.747141$ | $b=1.120839$ |
|  |  |  | $370$ | 370 | 370 | 370 | 370 |
|  |  | $0.001$ | 367.767 | 365.602 | 265.607 | 235.720 | 224.244 |
|  |  | 0.003 | 363.227 | 356.894 | 169.783 | 136.733 | 125.634 |
|  |  | 0.005 | 358.761 | 348.426 | 124.787 | 96.410 | 87.407 |
|  |  | 0.007 | 354.367 | 340.192 | 98.655 | 74.522 | 67.102 |
|  |  | 0.01 | 347.907 | 328.261 | 75.086 | 55.662 | 49.849 |
|  |  | 0.03 | 308.582 | 260.106 | 29.033 | 21.019 | 18.730 |
|  |  | 0.05 | 274.934 | 207.926 | 18.057 | 13.159 | 11.769 |
|  |  | 0.07 | 245.998 | 167.609 | 13.145 | 9.685 | 8.703 |
|  |  | 0.10 | 209.765 | 123.107 | 9.384 | 7.039 | 6.372 |
|  |  | 0.30 | 87.380 | 22.923 | 3.480 | 2.884 | 2.708 |
|  |  | 0.50 | 45.994 | 7.112 | 2.350 | 2.068 | 1.983 |
|  | RMI |  | 13.864 | 7.608 | 0.373 | 0.086 | 0.000 |
|  | -0.3 |  | $a=5, b=3.8159$ | $b=0.00032407$ | $b=0.2035172$ | $b=0.4066402$ | $b=0.609968$ |
|  |  |  | 370 | 370 | 370 | 370 | 370 |
|  |  | $0.001$ | 367.599 | 365.406 | 250.711 | 218.797 | 206.830 |
|  |  | 0.003 | 362.787 | 356.258 | 152.411 | 120.553 | 110.105 |
|  |  | 0.005 | 358.056 | 347.374 | 109.489 | 83.303 | 75.158 |
|  |  | 0.007 | 353.405 | 338.745 | 85.435 | 63.700 | 57.129 |
|  |  | 0.01 | 346.575 | 326.263 | 64.264 | 47.146 | 42.092 |
|  |  | 0.03 | 305.176 | 255.448 | 24.272 | 17.519 | 15.602 |
|  |  | 0.05 | 270.024 | 201.883 | 15.003 | 10.936 | 9.788 |
|  |  | 0.07 | 240.019 | 160.969 | 10.890 | 8.045 | 7.241 |
|  |  | 0.10 | 202.776 | 116.413 | 7.758 | 5.852 | 5.311 |
|  |  | 0.30 | 80.684 | 20.031 | 2.900 | 2.439 | 2.304 |
|  |  | 0.50 | 41.449 | 5.988 | 1.994 | 1.786 | 1.722 |
|  | RMI |  | 15.902 | 8.802 | 0.379 | 0.087 | 0.000 |



Fig. 2. The ARL for an $\operatorname{ARMAX}(2,2,2)$ process running on CUSUM, standard, and modified EWMA control charts. (a) The ARL for $\lambda=0.05$ where $\hat{\theta}_{1}=0.3$ and (b) the ARL for $\lambda=0.1$ where $\hat{\theta}_{1}=-0.3$.

According to Tables III and IV and Figs. 1 and 2, it is evident from the results that the ARL values for the explicit formulas method on the modified EWMA control chart were lower than those for the standard EWMA and CUSUM control charts for all $\lambda$, shift sizes and values of constants k , and thus its RMI values were lower.

## VII. Practical Applications with Real Data

We applied the explicit formulas for the ARL of an ARMAX $(1,1,1)$ process using 72 real data observations of the price of gasoline (Unit: USD per barrel [25]) and crude oil (Unit: USD per gallon [26]) from January 2015 to December 2020, with the latter being the explanatory variable, on CUSUM, and standard and modified EWMA control charts. The parameters were set as $\lambda=0.05$; the various parameter values listed in Tables V and VI ; and a shift size of $0.001,0.003,0.005,0.007,0.01,0.03,0.05$, $0.07,0.1,0.3$, or 0.5 . The results are summarized in Table VII.

We also performed another comparison for the ARL of an $\operatorname{ARMAX}(1,2,1)$ process involving 72 real-world data
observations of the exchange rate of 100 JPY to THB from January 2015 to December 2020, and with the USD to THB exchange rate over the same time period as the explanatory variable [27] on CUSUM, and standard and modified EWMA control charts. The parameters were set as $\lambda=0.05$; the various parameter values listed in Tables V and VI; and the same shift size as for the $\operatorname{ARMAX}(1,1,1)$ process. The results are summarized in Table VIII.

From Tables VII and VIII, it can be seen that the ARL values obtained from the explicit formulas running on the modified EWMA control chart were less than those for the CUSUM and standard EWMA control charts for all shift sizes and all values of $k$. Furthermore, as $k$ increased, ARL ${ }_{1}$ and the RMI decreased. Because the ARL of the CUSUM control chart is very different from EWMA and modified EWMA control chart, we compared the detection of shifts in the process means for the $\operatorname{ARMAX}(1,1,1)$ and ARMAX $(1,2,1)$ processes with real data on the two types of EWMA control charts only, the results for which are displayed in Figs. 3 and 4, respectively.

TABLE V
Fitting Statistics for the Real-World Datasets to ARMAX ( $1,1,1$ ) and ARMAX $(1,2,1)$ mOdels.

| Data | Variable | CoEFFICIENT | Std. Error | t | Sig. |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Gasoline (ARMAX (1,1,1)) | Constant ( $\hat{\omega}$ ) | 0.2345 | 0.0845 | 2.7740 | 0.0072 |
|  | $\operatorname{AR}(1)(\hat{\phi})$ | 0.5347 | 0.1367 | 3.9107 | 0.0002 |
|  | MA(1) ( $\hat{\theta}$ ) | 0.3361 | 0.1535 | 2.1904 | 0.0320 |
|  | Crude Oil ( $\hat{\beta}$ ) | 0.0267 | 0.0016 | 16.7076 | 0.0000 |
| JPY (ARMAX (1,2,1)) | Constant ( $\hat{\omega}$ ) | 15.4721 | 4.4510 | 3.4761 | 0.0009 |
|  | $\operatorname{AR}(1)(\hat{\phi})$ | 0.8267 | 0.0783 | 10.5550 | 0.0000 |
|  | $\operatorname{MA}(1)\left(\hat{\theta}_{1}\right)$ | 0.3982 | 0.1323 | 3.0107 | 0.0037 |
|  | $\operatorname{MA}(2)\left(\hat{\theta}_{2}\right)$ | 0.2744 | 0.1306 | 2.1009 | 0.0395 |
|  | $\operatorname{USD}(\hat{\beta})$ | 0.4365 | 0.1348 | 3.2378 | 0.0019 |

TABLE VI
Checking that Exponential Distributions Fit the White Noise of the Real-World
DATASETS

| DATASETS |  |  |  |
| :---: | :---: | :---: | :---: |
| Data | Mean $\left(\alpha_{0}\right)$ | Kolmogorov-Smirnov Z | Sig. |
| Gasoline (ARMAX(1,1,1)) | 0.0567 | 0.8942 | 0.4008 |
| JPY (ARMAX(1,2,1)) | 0.3331 | 0.4560 | 0.9854 |

IAENG International Journal of Applied Mathematics, 53:1, IJAM_53_1_32

TABLE VII
COMPARISON OF THE ARL FOR THE ARMAX $(1,1,1)$ PROCESS FOR REAL DATA RUNNING ON CUSUM, STANDARD, AND MODIFIED EWMA CONTROL Charts.

| $\lambda$ | $\delta$ | CUSUM | EWMA | Modified EWMA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | $k=1$ | $k=2$ | $k=3$ |
| 0.05 |  | $a=60, b=45.458$ | $b=1.989 \times 10^{-16}$ | $b=0.021604$ | $b=0.044232$ | $b=0.0668341$ |
|  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  | 0.001 | 367.811 | 357.104 | 313.520 | 267.902 | 247.382 |
|  | 0.003 | 363.379 | 331.803 | 239.977 | 172.499 | 148.884 |
|  | 0.005 | 359.017 | 308.387 | 194.209 | 127.224 | 106.576 |
|  | 0.007 | 354.725 | 286.710 | 162.982 | 100.788 | 83.047 |
|  | 0.01 | 348.413 | 256.956 | 131.168 | 76.855 | 62.445 |
|  | 0.03 | 309.928 | 126.171 | 56.204 | 29.841 | 23.787 |
|  | 0.05 | 276.914 | 63.833 | 35.256 | 18.587 | 14.867 |
|  | 0.07 | 248.452 | 33.382 | 25.446 | 13.543 | 10.907 |
|  | 0.10 | 212.704 | 13.500 | 17.761 | 9.676 | 7.885 |
|  | 0.30 | 90.643 | 1.067 | 5.565 | 3.595 | 3.132 |
|  | 0.50 | 48.538 | 1.001 | 3.314 | 2.423 | 2.202 |
| RMI |  | 18.415 | 1.625 | 1.301 | 0.470 | 0.261 |

TABLE VIII
COMPARISON OF THE ARL FOR THE ARMAX(1,2,1) PROCESS FOR REAL DATA RUNNING ON CUSUM, STANDARD, AND MODIFIED EWMA CONTROL CHARTS.

| $\lambda$ | $\delta$ | CUSUM | EWMA | Modified EWMA |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | $k=1$ | $k=2$ | $k=3$ |
| 0.05 |  | $a=10, b=8.6369$ | $b=3.975 \times 10^{-13}$ | $b=0.231849$ | $b=0.470822$ | $b=0.709694$ |
|  | 0.00 | 370 | 370 | 370 | 370 | 370 |
|  | 0.001 | 367.750 | 358.642 | 313.296 | 275.493 | 259.122 |
|  | 0.003 | 363.261 | 336.976 | 239.605 | 182.407 | 162.178 |
|  | 0.005 | 358.844 | 316.701 | 193.887 | 136.406 | 118.143 |
|  | 0.007 | 354.498 | 297.720 | 162.756 | 108.976 | 92.989 |
|  | 0.01 | 348.109 | 271.490 | 131.089 | 83.770 | 70.560 |
|  | 0.03 | 309.185 | 148.934 | 56.649 | 33.183 | 27.415 |
|  | 0.05 | 275.844 | 83.758 | 35.868 | 20.849 | 17.252 |
|  | 0.07 | 247.141 | 48.296 | 26.123 | 15.288 | 12.712 |
|  | 0.10 | 211.152 | 22.215 | 18.468 | 11.006 | 9.232 |
|  | 0.30 | 89.012 | 1.257 | 6.149 | 4.191 | 3.703 |
|  | 0.50 | 47.332 | 1.010 | 3.771 | 2.836 | 2.591 |
| RMI |  | 15.977 | 1.724 | 1.147 | 0.472 | 0.293 |



Fig. 3. Mean shift detection for the $\operatorname{ARMAX}(1,1,1)$ process for the price of gasoline and crude oil is explanatory variable.

The results in Fig. 3 display that the modified EWMA control chart was able to detect a change in the price of gasoline for the first time at the $5^{\text {th }}$ observation while the
standard EWMA control chart achieved this at the $45^{\text {th }}$ observation.

The results in Fig. 4 indicate that the modified EWMA control chart can be detect a change in the exchange rate of


Fig. 4. Mean shift detection for the ARMAX $(1,2,1)$ process for the exchange rate of Japanese Yen and US Dollars is explanatory variable.

JPY for the first time at the $16^{\text {th }}$ observation while the standard EWMA control chart achieved this at the $22^{\text {nd }}$ observation. Hence, in both cases, the performance of the modified EWMA control chart is better than of the standard EWMA control chart for detecting shifts in the process mean, thus the former is more efficient than the latter.

## VIII. Conclusions

We derived explicit formulas for the ARL of the modified EWMA control chart for an ARMAX(p,q,r) process. and used simulated data to check its accuracy by comparing it with the ARL derived from the NIE method by using an absolute percentage difference. The results indicate that although both methods yielded very close ARL values with an absolute percentage difference of less than $0.00001 \%$, the explicit formula method took much less time to calculate them. A comparison of the ARL derived by using explicit formulas for the ARL of an ARMAX(p,q,r) process with exponential white noise running on CUSUM, and standard and modified EWMA control charts, indicate that the proposed explicit formulas was more effective than on the CUSUM and standard EWMA control charts in terms of RMI. Practical application with real data for ARMAX(p,q,r) processes with exponential white noise running on the three control charts indicate that the method on the modified EWMA control chart performed much better than on the other two for a one-sided shift. In addition, as $k$ increased, its ARL $_{1}$ and the RMI decreased. Based on the findings, the ARL derived by using explicit formulas of an ARMAX( $\mathrm{p}, \mathrm{q}, \mathrm{r}$ ) process with exponential white noise running on the modified EWMA control chart were the most efficient.

## REFERENCES

[1] A. Koetsier, S. N. van der Veer, K. J. Jager, N. Peek, and N. F. de Keizer. (2012, April). Control charts in healthcare quality improvement. A systematic review on adherence to methodological criteria. Methods Inf Med. (online). 51(3) pp. 189-198. Available: doi: 10.3414/ME11-01-0055
[2] R. Poovarasan, S. Keerthi, K. Yuvashree, and C. Thirumalai. (2017, May). " Analysis on diabetes patients using Pearson, cost optimization, control chart" . In Proc. International Conference on Trends in Electronics and Informatics, Tirunelveli, 2017, pp. 11391142. Available: doi: 10.1109/ICOEI.2017.8300891
[3] M. Kovarik, and P. Klimek. (2012, September). The usage of time series control charts for financial process analysis. Journal of Competitiveness (online). 4(3): 29-45. Available: https://doi.org/10.7441/joc.2012.03.03
[4] W. A. Shewhart. Economic control of quality of manufactured product, D. Van Nostrand Company, New York, NY, 1931.
[5] E. S. Page. (1954, June). Continuous inspection schemes. Biometrika. (online). 41(1/2): 100-115. Available: doi:10.1093/BIOMET/41.1-2.100
[6] S. W. Roberts. (1959, August). Control chart test based on geometric moving average. Technometrics. (online). 1(3): 239-250. Available: doi: 10.1080/00401706.1959.10489860
[7] N. Khan, M. Aslam, and C-H. Jun. (2016, October). Design of a control chart using a modified EWMA statistic. Quality and Reliability Engineering International. (online). 33: 1095-1104. Available: https://doi.org/10.1002/qre. 2102
[8] A. K. Patel, and J. Divecha. (2011, January). Modified exponentially weighted moving average (EWMA) control chart for an analytical process data. Journal of Chemical Engineering and Materials Science. (online). 2(1): 12-20. Available: http://www.academicjournals.org/journal/JCEMS/article-full-textpdf/466796E1469
[9] D. C. Montgomery. Introduction to Statistical Quality Control 7th ed., John Wiley and Sons Inc., Hoboken, NJ, 2012.
[10] Y. Areepong, "An integral equation approach for analysis of control charts," Ph.D. thesis, Dept. Mathematical Science. Eng, University of Technology, Sydney, Sydney, Australia, 2009 (online). Available:
https://www.semanticscholar.org/paper/An-integral-equation-approach-for-analysis-of-
Areepong/890820563ce06bd528c8f5d676cce45e6efad507
[11] S. V. Crowder. (1987, November). A simple method for studying run length distributions of exponentially weighted moving average charts. Technometrics. (online). 29(4): 401-407. Available: http://jstor.org/stable/1269450
[12] C. W. Champ, and S. E. Rigdon. (1991, January). A comparison of the Markov chain and the integral equation approaches for evaluating the run length distribution of quality control charts. Communications in Statistics-Simulation and Computation. (online). 20(1): 191-204. Available: https://doi.org/10.1080/03610919108812948
[13] Y. Areepong, and A. Novikov. (2008, July). Martingale approach to EWMA control chart for changes in exponential distribution. Journal of Quality Measurement and Analysis. (online). 4(1): 197-203. Available: http://journalarticle.ukm.my/1867
[14] S. Phanyaem, Y. Areepong, and S. Sukparungsee. (2014, July). Explicit formulas of average run length for $\operatorname{ARMA}(1,1)$ process of CUSUM control chart. Far East Journal of Mathematical Sciences. (online). 211-224. Avaliable: http://www.pphmj.com/abstract/8583.htm
[15] S. Sukparungsee, and Y. Areepong. (2017, July). An explicit analytical solution of the average run length of an exponentially weighted moving average control chart using an autoregressive model. Chiang Mai Journal of Science. (online). 44(3): 1172-1179. Available: https://epg.science.cmu.ac.th/ejournal/journaldetail.php?id=8305
[16] R. Sunthornwat, Y. Areepong, and S. Sukparungsee. (2018, July). Average run length with a practical investigation of estimating parameters of the EWMA control chart on the long memory AFRIMA process. Thailand Statistician. (online). 16(2): 190-202. Available: https://ph02.tci-thaijo.org/index.php/thaistat/article/view/135562
[17] W. Peerajit, and Y. Areepong. (2022, January). The performance of CUSUM control chart for monitoring process mean for autoregressive moving average with exogenous variable model. Applied Science and Engineering Progress. (online). 15(1): 1-10. Available: https://doi.org/10.14416/j.asep.2020.11.007
[18] Y. Supharakonsakun, Y. Areepong, and S. Sukparungsee. (2020, January). The exact solution of the average run length on a modified EWMA control chart for the first-order moving-average process. ScienceAsia. (online). 46: 109-118. Available: doi: 10.2306/scienceasia1513-1874.2020.015
[19] P. Phanthuna, Y. Areepong, and S. Sukparungsee. (2021, August). Run length distribution for a modified EWMA scheme fitted with a stationary AR(p) model. Communications in Statistics - Simulation and Computation. (online). 1-20. Available: doi: 10.1080/03610918.2021.1958847
[20] K. Silpakob, Y. Areepong, and S. Sukparungsee, and R. Sunthornwat. (2021, September). Explicit analytical solutions for the average run length of modified EWMA control chart for ARX(p,r) processes. Songklanakarin Journal of Science and Technology. (online). 43(5): 1414-1427. Available: doi: $10.14456 /$ sjst-psu. 2021.185
[21] P. Phanthuna, and Y. Areepong. (2022, October). Detection sensitivity of a modified EWMA control chart with a time series model with fractionality and integration. Emerging Science Journal. (noline). 6(5):1134-1152. Avaliable: https://ijournalse.org/index.php/ESJ/article/view/1159
[22] G. Mititelu, Y. Areepong, S. Sukparungsee, and A. Novikov. (2010, January). Explicit analytical solutions for the average run length of CUSUM and EWMA charts. Contribution in Mathematics and Applications III East-West J. of Mathematics. (online). special volume: 253-265. Available: https://www.researchgate.net/publication/266832198_Explicit_analyti cal_solutions_for_the_average_run_length_of_CUSUM_and_EWMA _charts
[23] S. P. Richard. (2007, November). A simple proof of the Banach contraction principle. Journal of Fixed Point Theory and Applications. (online). 2: 221-223. Available: doi: 10.1007/s11784-007-0041-6
[24] P. Phanthuna, Y. Areepong, and S. Sukparungsee. "Numerical integral equation methods of average run length on modified EWMA control chart for exponential $\operatorname{AR}(1)$ process," Lecture Notes in Engineering and Computer Science: Proceedings of The International MultiConference of Engineers and Computer Scientists 2018, 14-16 March, 2018, Hong Kong, pp845-847.
[25] IndexMundi. The price of gasoline. 2022. Retrieved from https://www.indexmundi.com/commodities/?commodity=gasoline\&m onths=120
[26] IndexMundi. The price of gasoline. 2022. Retrieved from https://www.indexmundi.com/commodities/?commodity=crude-oil-west-texas-intermediate\&months $=120$
[27] Bank of Thailand. Exchange rates of commercial banks in Bangkok. 2022. Retrieved from https://www.bot.or.th/App/BTWS_STAT/statistics/ReportPage.aspx?r eportID=123\&language=th


[^0]:    Manuscript received July 05, 2022; revised January 13, 2023.
    This research was supported by the Thailand Science Research and Innovation Fund, and King Mongkut's University of Technology North Bangkok Contract no. KMUTNB-65-BASIC-15.

    Korakoch Silpakob is a PhD student of Department of Applied Statistics, Faculty of Applied Science, King Mongkut's University of Technology North Bangkok, Bang Sue, Bangkok, 10800, Thailand (e-mail: korakoch14737@gmail.com).

    Yupaporn Areepong is a professor of Department of Applied Statistics, Faculty of Applied Science, King Mongkut's University of Technology North Bangkok, Bang Sue, Bangkok, 10800, Thailand (Corresponding author to provide e-mail: yupaporn.a@sci.kmutnb.ac.th).

    Saowanit Sukparungsee is a professor of Department of Applied Statistics, Faculty of Applied Science, King Mongkut's University of Technology North Bangkok, Bang Sue, Bangkok, 10800, Thailand (e-mail: swns@kmutnb.ac.th).

    Rapin Sunthornwat is an assistant professor of Industrial Technology Program, Faculty of Science and Technology, Pathumwan Institute of Technology, Pathumwan, Bangkok, 10330, Thailand (e-mail: rapin@pit.ac.th).

