# An Alternative Numerical Method for Finding the Value of Transformation Parameter in the Box-Cox Transformation

### Lakhana Watthanacheewakul

Abstract—To analyze non normal data, the data should be transformed to a normal distribution. The well-known Box-Cox transformation can be used to get the normality. Since the transformation parameter is usually unknown, many statisticians studied the methods of estimation of the transformation parameter,  $\lambda$ , appearing in the Box-Cox transformation. Moreover, we can use the numerical methods for single variable function optimization to find the value of the transformation parameter  $\lambda$  such as bisection method, Newton's method, secant method, and others. The alternative numerical method  $\lambda$  is applied to find the suitable value of  $\lambda$ . It is better than the bisection method in the context of the number of function evaluations and the processing time.

*Index Terms*— Alternative numerical method, Bisection method, Box-Cox transformation, Transformation parameter

## I. INTRODUCTION

In statistical data analysis the basic assumptions are that the model is additive and the errors are randomly, independently, and normally distributed about zero mean and constant variance. If the basic assumptions are not satisfied, then the statistical inference under the normal theory cannot be applied. Tukey [1] suggested that in analyzing data which do not match the assumptions of the conventional method of analysis, we have two alternative ways to go about. We may transform the data to fit the assumptions, or we may develop some new methods of analysis with assumptions fitting the original data. If we can find a satisfactory transformation, it will almost always be easier to use the conventional method of analysis rather than to develop a new one. The well-known Box-Cox transformation is often used to transform non normal data in regression analysis and the analysis of variance to normality with homogeneity of variances. Most statisticians considered the problem of estimation of the transformation parameter,  $\lambda$ , in the Box-Cox transformation [2-5]. Box and Cox [6] proposed the maximum likelihood method to obtain the proper value of  $\lambda$  by plotting the log likelihood against  $\lambda$  for a trial series of values. From this plot the maximizing value  $\hat{\lambda}$  may be read off. Moreover, we can

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use the numerical methods for single variable function optimization to find the value of transformation parameter  $\lambda$  such as bisection method, Newton's method, secant method, and others [7-9]. In this paper, the alternative numerical method for finding the appropriate value of  $\lambda$  is applied for k exponential populations with one parameter.

## II. ESTIMATION OF THE TRANSFORMATION PARAMETER

The Box-Cox transformation for any sets of exponential data to normality

$$Y_{ij} = \begin{cases} \frac{X_{ij}^{\lambda} - 1}{\lambda}, & \lambda \neq 0\\ \ln X_{ij}, & \lambda = 0, \end{cases}$$
(1)

for  $x_{ij} > 0$ , where  $X_{ij}$  is a random variable in the *jth* trial from the *ith* population,  $Y_{ij}$  the transformed variable of  $X_{ij}$  and  $\lambda$ a transformation parameter. The probability density function of each  $Y_{ij}$  is assumed to be normal with mean  $\mu_i$  and variance  $\sigma^2$ , i.e.

$$f(y_{ij}|\mu_{i},\sigma^{2}) = \frac{1}{\sqrt{2\pi\sigma^{2}}} \exp\left\{\frac{-(y_{ij}-\mu_{i})^{2}}{2\sigma^{2}}\right\}, \quad (2)$$

where  $y_{ij}$  is the observed value of  $Y_{ij}$ .

The likelihood function in terms of  $\mu_i$  and  $\sigma^2$  of  $Y_{ij}$  ,  $i=1,...,k;\,j=1,...,n_i$ 

$$L(\mu_{i},\sigma^{2}|y_{ij}) = \frac{1}{(2\pi\sigma^{2})^{n/2}} \exp\left\{-\frac{1}{2\sigma^{2}}\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}(y_{ij}-\mu_{i})^{2}\right\}.$$
 (3)

The maximum likelihood estimators of  $\mu_i$  and  $\sigma^2$  are

$$\hat{\mu}_i = \frac{1}{n_i} \sum_{j=1}^{n_i} Y_{ij}$$
 and  $\partial^2 = \frac{1}{n} \sum_{i=1}^k \sum_{j=1}^{n_i} (Y_{ij} - \mu_i)^2$ , respectively.

The likelihood in relation to the sample observations is given by

$$L(\mu_i, \sigma^2 | x_{ij}) =$$

$$\frac{1}{\left(2\pi\sigma^{2}\right)^{n/2}}\exp\left\{-\frac{1}{2\sigma^{2}}\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}\left[\frac{x_{ij}^{\lambda}-1}{\lambda}-\mu_{i}\right]^{2}\right\}J(\lambda;x)$$

$$(4)$$
where
$$J(\lambda;x) = \prod_{i=1}^{k}\prod_{j=1}^{n_{i}}\left|\frac{\partial y_{ij}}{\partial x_{ij}}\right| \qquad = \prod_{i=1}^{k}\prod_{j=1}^{n_{i}}\left|\frac{\partial}{\partial x_{ij}}\left[\frac{x_{ij}^{\lambda}-1}{\lambda}\right]\right|$$

$$= \prod_{i=1}^{k}\prod_{j=1}^{n_{i}}x_{ij}^{\lambda-1}.$$

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Thus, the log likelihood of the observations is

$$\ln L(\mu_{i}, \sigma^{2}, \lambda | \mathbf{x}_{ij}) = -\frac{n}{2} \ln 2\pi - \frac{n}{2} \ln \sigma^{2} - \frac{1}{2\sigma^{2}} \sum_{i=1}^{k} \sum_{j=1}^{n_{i}} \left[ \frac{\mathbf{x}_{ij}^{\lambda} - 1}{\lambda} - \mu_{i} \right]^{2} + (\lambda - 1) \sum_{i=1}^{k} \sum_{j=1}^{n_{i}} \ln \mathbf{x}_{ij}.$$
(5)

The log likelihood function of  $\lambda$  is, except for a constant,

$$\begin{split} \ln L(\lambda | \mathbf{x}_{ij}) &= -\frac{n}{2} \ln \left( \frac{1}{n} \sum_{i=1}^{k} \sum_{j=1}^{n_i} \left\{ \frac{\mathbf{x}_{ij}^{\lambda} - 1}{\lambda} - \frac{1}{n_i} \sum_{j=1}^{n_i} \left[ \frac{\mathbf{x}_{ij}^{\lambda} - 1}{\lambda} \right] \right\}^2 \right) \\ &+ (\lambda - 1) \sum_{i=1}^{k} \sum_{j=1}^{n_i} \ln \mathbf{x}_{ij}. \end{split}$$
(6)

The partial derivative of  $\ln L(\lambda \big| x_{ij})$  with respect to  $\lambda$  is  $d \ln L(\lambda)$ 

$$\frac{d\lambda}{-n\left[\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}x_{ij}^{2\lambda}\ln x_{ij} - \sum_{i=1}^{k}\frac{1}{n_{i}}\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\right)\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\ln x_{ij}\right)\right]}{\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}x_{ij}^{2\lambda} - \sum_{i=1}^{k}\frac{1}{n_{i}}\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\right)^{2}}$$
(7)  
$$+\frac{n}{\lambda} + \sum_{i=1}^{k}\sum_{j=1}^{n_{i}}\ln x_{ij}.$$

The maximum likelihood estimate of  $\lambda$  is obtained by solving the likelihood equation

$$\frac{-n\left[\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}x_{ij}^{2\lambda}\ln x_{ij} - \sum_{i=1}^{k}\frac{1}{n_{i}}\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\right)\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\ln x_{ij}\right)\right]}{\sum_{i=1}^{k}\sum_{j=1}^{n_{i}}x_{ij}^{2\lambda} - \sum_{i=1}^{k}\frac{1}{n_{i}}\left(\sum_{j=1}^{n_{i}}x_{ij}^{\lambda}\right)^{2}}$$

$$+\frac{n}{\lambda} + \sum_{i=1}^{k}\sum_{j=1}^{n_{i}}\ln x_{ij} = 0.$$
(8)

The maximum likelihood estimate of  $\lambda$  is obtained by solving the likelihood equation. This equation is not easy to solve for  $\lambda$  in terms of the observations  $x_{ij}$  and, hence, a numerical approach should be considered.

#### III. AN ALTERNATIVE NUMERICAL METHOD

Since the curvature of the log likelihood function is of a quadratic pattern, it has only one maximum value. If it was ever desired to determine  $\hat{\lambda}$  more precisely this could be done by numerically determining the value  $\hat{\lambda}$  for which the derivatives with respect to  $\lambda$  are all zero [6]. Numerical methods, such as the bisection method, can be applied, but here an alternative numerical method is applied to find  $\lambda$ . The idea behind this method is that the slope of curvature of the log likelihood function is used to find  $\lambda$  and the transformation parameter is obtained when the slope is nearly zero. The first derivative of the log likelihood function of  $\lambda$ , or the slope given by (7), is denoted by  $f'(\lambda)$ . In what follows it will always be assumed that  $f'(\lambda)$  is continuous on  $[\lambda_1, \lambda_2]$ . Watthanacheewakul [10] proposed the algorithm for the alternative numerical method to obtain  $\hat{\lambda}$  in the alternative transformation is as follows:

## Step 1 The Appropriate Interval

**Step 1.1** Initialize  $\lambda_1, \lambda_2$ , and  $\epsilon$ 

First of all, the tolerance value for the stopping rule,  $\varepsilon$ , is set at 0.0001 and  $\lambda_1, \lambda_2$  are arbitrary values where  $\lambda_1 < \lambda_2$ .

Step 1.2 Calculate the value of the first derivatives of the log likelihood of  $\lambda_1$  and  $\lambda_2$ 

The values of the first derivatives of the log likelihood of  $\lambda_1$  and  $\lambda_2$  are calculated by replacing  $\lambda$  in  $f'(\lambda)$  with  $\lambda_1$  and  $\lambda_2$  respectively.

**Step 1.3** Compare  $|f'(\lambda_1)|$  and  $|f'(\lambda_2)|$  with  $\varepsilon$ 

If  $|f'(\lambda_1)| < \varepsilon$  then let  $\lambda = \lambda_1$ . If  $|f'(\lambda_1)| \ge \varepsilon$ , then consider  $|f'(\lambda_2)|$ . If  $|f'(\lambda_2)| < \varepsilon$ , then  $\lambda = \lambda_2$ . Otherwise, proceed to **Step 1.4**.

**Step 1.4** Find the appropriate interval  $[\lambda_1, \lambda_2]$ 

If  $|f'(\lambda_1)| \ge \varepsilon$  and  $f'(\lambda_1) < 0$ , let  $m = |\lambda_1 - \lambda_2|$ , then set  $\lambda'_2 = \lambda_1$  and  $\lambda'_1 = \lambda'_2 - \frac{m}{4} = \lambda_1 - \frac{m}{4}$ . Set  $\lambda_1 = \lambda'_1$ . Repeat this process until  $f'(\lambda_1) > 0$ .

If  $|f'(\lambda_2)| \ge \varepsilon$  and  $f'(\lambda_2) > 0$ , let  $m = |\lambda_1 - \lambda_2|$ , then set  $\lambda'_1 = \lambda_2$  and  $\lambda'_2 = \lambda'_1 + m/_4 = \lambda_2 + m/_4$ . Set  $\lambda_2 = \lambda'_2$ . Repeat this process until  $f'(\lambda_2) < 0$ .

If  $f'(\lambda_1) > 0$  and  $f'(\lambda_2) < 0$  then proceed to **Step 2.** 

# **Step 2** The Optimization Value

**Step 2.1** Compare  $|f'(\lambda_1)|$  and  $|f'(\lambda_2)|$ 

 $\begin{array}{ll} \mbox{Calculate} & m = \left|\lambda_1 - \lambda_2\right| & . \quad \mbox{If} & \left|f'(\lambda_1)\right| < \left|f'(\lambda_2)\right| & , \quad \mbox{let} \\ \lambda = \lambda_1 + \frac{m_A'}{4} & , \quad \mbox{otherwise let} \ \lambda = \lambda_2 - \frac{m_A'}{4} & . \end{array}$ 

**Step 2.2** Check for a stopping rule

$$\begin{split} & If \left| f'(\lambda) \right| < \epsilon \text{ , stop the iteration. If } \left| f'(\lambda) \right| \geq \epsilon \text{ and } f'(\lambda) > 0 \text{ ,} \\ & \text{let } \lambda_1 = \lambda \text{ and go back to } \textbf{Step 2.1. If } \left| f'(\lambda) \right| \geq \epsilon \text{ and} \\ & f'(\lambda) < 0 \text{ , let } \lambda_2 = \lambda \text{ and go back to } \textbf{Step 2.1.} \end{split}$$

This algorithm is applied to find the value of transformation parameter in the Box-Cox transformation.

## IV. APPLYING THE ALTERNATIVE NUMERICAL METHOD

The random samples of size 20 were taken from exponential population with scale parameter ( $\beta$ ) equals three. The data are as follows 13.51945, 0.071361, 0.347777, 5.577101, 0.605019, 5.295623, 0.487435, 1.374586, 8.99852, 0.91593, 1.540253, 5.55445, 9.484076, 5.394369, 1.776799, 2.992351, 1.597922, 3.602009, 0.179283, 3.424097. They are used for finding the transformation parameter using the alternative numerical method. The procedure is as follows:

#### Step 1 The Appropriate Interval Round No.1

**Step 1.1** Let  $\lambda_1 = -1.0$ ,  $\lambda_2 = 0.01$  and  $\epsilon = 0.0001$ .

**Step 1.2** Calculate  $f'(\lambda_1) = 44.442623$ 

and  $f'(\lambda_2) = 9.038474$ .

**Step 1.3**  $|f'(\lambda_1)| \ge \varepsilon$  and  $|f'(\lambda_2)| \ge \varepsilon$ .

Step 1.4 Since 
$$f'(\lambda_2) > 0$$
,  $m = |\lambda_1 - \lambda_2| = |-1 - 0.01| = 1.01$ .  
Set  $\lambda_1' = 0.01$  and  $\lambda_2' = \lambda_1' + \frac{m}{4} = \lambda_2 + \frac{m}{4}$ 

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= 0.01+0.2525 = 0.2625. Set 
$$\lambda_2 = \lambda'_2$$
 and

 $f'(\lambda_2) = -0.553462.$ 

The appropriate interval is [0.01, 0.2625].

# Step 2 The optimization value

# Round No.1

Step 2.1 Calculate  $m = |\lambda_1 - \lambda_2| = |0.01 - 0.2625| = 0.2525$ .  $f'(\lambda_1) = 9.038474$  and  $f'(\lambda_2) = -0.553462$ .

Since 
$$|f'(\lambda_1)| > |f'(\lambda_2)|$$
,  $\lambda = \lambda_2 - \frac{m}{4} = 0.199375$ .

**Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = 1.673146$ , let

 $\lambda_1 = \lambda = 0.199375.$ 

The new interval is [0.199375, 0.262500].

## Round No.2 Step 2.1 Calculate

 $m = |\lambda_1 - \lambda_2| = |0.199375 - 0.262500| = 0.063125.$ f'(\lambda\_1) = 1.673146 and f'(\lambda\_2) = -0.553462.

Since  $|f'(\lambda_1)| > |f'(\lambda_2)|$ ,  $\lambda = \lambda_2 - m_A = 0.246719$ .

**Step2.2** Since 
$$|f'(\lambda)| \ge \varepsilon$$
 and  $f'(\lambda) = -0.008474$ , let

 $\lambda_2 = \lambda = 0.246719.$ 

The new interval is [0.199375, 0.246719].

## Round No.3

Step 2.1 Calculate  $m = |\lambda_1 - \lambda_2| = |0.199375 \cdot 0.246719| = 0.047344.$ 

 $f'(\lambda_1) = 1.673146$  and  $f'(\lambda_2) = -0.008474$ .

Since  $|f'(\lambda_1)| > |f'(\lambda_2)|$ ,  $\lambda = \lambda_2 - \frac{m}{4} = 0.234883$ .

**Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = 0.405402$ , let

 $\lambda_1 = \lambda = 0.234883.$ 

The new interval is [0.234883, 0.246719].

Round No.4 Step 2.1 Calculate

 $m = \left| \lambda_1 - \lambda_2 \right| = \left| 0.234883 \cdot 0.246719 \right| = 0.011836 \,.$ 

 $f'(\lambda_1) = 0.405402$  and  $f'(\lambda_2) = -0.008474$ .

Since 
$$|f'(\lambda_1)| > |f'(\lambda_2)|$$
,  $\lambda = \lambda_2 - \frac{m}{4} = 0.243760$ .

**Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = 0.094584$ , let

 $\lambda_1 = \lambda = 0.243760.$ 

The new interval is [0.243760, 0.246719].

# Round No.5

 $m = |\lambda_1 - \lambda_2| = |0.243760 - 0.246719| = 0.002959.$ f'(\lambda\_1) = 0.094584 and f'(\lambda\_2) = -0.008474.

Since 
$$|\mathbf{f}'(\lambda_1)| > |\mathbf{f}'(\lambda_2)|$$
,  $\lambda = \lambda_2 - m_A = 0.245979$ 

**Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = 0.017265$ , let

 $\lambda_1 = \lambda = 0.245979.$ 

The new interval is [0.245979, 0.246719].

## **Round No.6**

Step 2.1 Calculate

$$\begin{split} m &= \left|\lambda_1 - \lambda_2\right| = \left|0.245979 - 0.246719\right| = 0.000740 \ . \\ f'(\lambda_1) &= \ 0.017265 \ \text{and} \ f'(\lambda_2) = -0.008474. \end{split}$$

Since  $|\mathbf{f}'(\lambda_1)| > |\mathbf{f}'(\lambda_2)|$ ,  $\lambda = \lambda_2 - m_{\underline{\lambda}} = 0.246534$ . **Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = -0.002040$ , let  $\lambda_2 = \lambda = 0.246534.$ The new interval is [0.245979, 0.246534]. Round No.7 Step 2.1 Calculate  $m = |\lambda_1 - \lambda_2| = |0.245979 - 0.246534| = 0.000555$ .  $f'(\lambda_1) = 0.017265$  and  $f'(\lambda_2) = -0.002040$ . Since  $|f'(\lambda_1)| > |f'(\lambda_2)|$ ,  $\lambda = \lambda_2 - m_A = 0.246395$ . Step 2.2 Since  $|f'(\lambda)| \ge \epsilon$  and  $f'(\lambda) = 0.002785$ , let  $\lambda_1 = \lambda = 0.246395.$ The new interval is [0.246395, 0.246534]. Round No.8 Step 2.1 Calculate  $m = |\lambda_1 - \lambda_2| = |0.246395 - 0.246534| = 0.000139$ .  $f'(\lambda_1) = 0.002785$  and  $f'(\lambda_2) = -0.002040$ . Since  $|f'(\lambda_1)| > |f'(\lambda_2)|$ ,  $\lambda = \lambda_2 - m_A = 0.246499$ . **Step 2.2** Since  $|f'(\lambda)| \ge \varepsilon$  and  $f'(\lambda) = -0.000834$ , let  $\lambda_2 = \lambda = 0.246499.$ The new interval is [0.246395, 0.246499]. Round No.9 Step 2.1 Calculate  $m = |\lambda_1 - \lambda_2| = |0.246395 - 0.246499| = 0.000104$ .  $f'(\lambda_1) = 0.002785$  and  $f'(\lambda_2) = -0.000834$ . Since  $|f'(\lambda_1)| > |f'(\lambda_2)|$ ,  $\lambda = \lambda_2 - m_{1/2} = 0.246473$ . **Step 2.2** Since  $|f'(\lambda)| = 0.000071$ ,  $|f'(\lambda)| < \varepsilon$ . So  $\lambda = 0.246473.$ The value of transformation parameter is  $\lambda = 0.246473$ .

# V. THE SIMULATION METHOD

Exponential populations of size  $N_i = 5,000$  were generated for the values of parameters  $\beta_i$ , for i = 1,...,k. From each generated population, 1,000 random samples, each of size  $n_i$ , for i = 1,...,k, were drawn.

The values of parameters are set as follows:

1) k = number of the populations = 3

2)  $n_i$  = sample sizes from the *ith* exponential population between 10 and 150, for i = 1, ..., k

3)  $\beta_i$ , the scale parameter of the *ith* exponential population, is between 1 and 5.

The results are shown in Table I-III. For any k exponential populations, both equal sample sizes and unequal sample sizes, both equal scale parameters and unequal scale parameters, the alternative numerical method for finding the transformation parameter,  $\lambda$ , is better than the bisection

method in the context of the number of function evaluations and processing time.

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Table I	Average of Number of Function Evaluations and
	Processing Time for the Bisection and Alternative
	Numerical Methods for k=1

n	β	Average of Number of Function Evaluations		Processing Time (Seconds)		
		Bisection	Alternative	Bisection	Alternative	
10	1	16.890	12.283	5.250	4.937	
	3	16.873	12.255	5.250	5.062	
	5	17.115	12.405	5.093	4.859	
30	1	19.112	13.495	6.000	5.515	
	3	18.913	13.515	6.235	5.672	
	5	19.131	13.767	6.000	5.687	
100	1	21.030	15.127	9.422	8.578	
	3	21.326	15.378	10.141	8.766	
	5	21.078	15.144	10.140	8.656	

Table IIAverage of Number of Function Evaluations and<br/>Processing Time for the Bisection and Alternative<br/>Numerical Methods for k=2

n <sub>i</sub>	Scale parameter		Average of Number of Function Evaluations		Processing Time (Seconds)	
	β	$\beta_2$	Bisection	Alternative	Bisection	Alternative
$n_1 = 10,$	1	1	18.290	13.167	5.593	5.328
n <sub>2</sub> = 10	3	5	18.776	13.795	5.844	5.594
$n_1 = 10,$	1	1	18.685	13.463	5.921	5.485
n <sub>2</sub> = 15	3	5	18.876	13.637	5.906	5.500
$n_1 = 30,$	1	1	20.428	14.675	7.718	6.906
$n_2 = 30$	3	5	20.420	14.627	8.063	7.141
$n_1 = 30,$	1	1	20.669	14.886	8.687	7.562
n <sub>2</sub> = 45	3	5	20.991	15.105	8.828	7.734
$n_1 = 100,$	1	1	22.111	16.061	16.297	13.422
n <sub>2</sub> = 100	3	5	22.459	16.242	17.547	13.594
$n_1 = 100,$ $n_2 = 150$	1	1	22.476	16.333	19.938	16.063
n <sub>2</sub> = 150	3	5	22.593	16.425	19.765	15.969

## VI. CONCLUSION

For analyzing data which do not match the assumptions of the conventional method of analysis, the data should first be transformed to fit all the assumptions required. The well-known Box-Cox transformation can be used to get the normality. The transformation parameter ( $\lambda$ ) is unknown. The alternative numerical method is applied to find the appropriate value of  $\lambda$ . The value of it is obtained when the slope of curvature of the log likelihood function is nearly zero. From simulation results, this method is better than the bisection method in the context of the number of function evaluations and the processing time.

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## Table III Average of Number of Function Evaluations and Processing Time for the Bisection and Alternative Numerical Methods for k=3

n <sub>i</sub>	Scale parameter			Average of Number of Function Evaluations		Processing Time (Seconds)	
	$\beta_1$	$\beta_2$	β <sub>3</sub>	Bisection	Alternative	Bisection	Alternative
$n_1 = 10$ ,	1	1	1	19.109	13.656	6.375	5.938
$n_2 = 10,$ $n_3 = 10$	1	3	5	19.952	14.645	6.234	5.828
$n_1 = 10$ ,	1	1	1	20.226	14.513	8.281	7.078
$n_2 = 20,$ $n_3 = 30$	1	3	5	20.558	14.756	8.234	7.328
$n_1 = 30$ ,	1	1	1	20.978	15.249	7.515	6.360
$n_2 = 30,$ $n_3 = 30$	1	3	5	21.501	15.414	7.593	6.407
$n_1 = 30,$	1	1	1	21.659	15.551	10.032	8.235
$n_2 = 45, n_3 = 60$	1	3	5	22.216	15.827	10.094	8.235
$n_1 = 100$ ,	1	1	1	22.739	16.491	18.860	14.812
$n_2 = 100,$ $n_3 = 100$	1	3	5	23.259	16.526	17.985	14.843
$n_1 = 100,$	1	1	1	23.519	17.095	27.750	21.484
$n_2 = 150,$ $n_3 = 200$	1	3	5	23.826	16.767	27.829	21.015